

The effect of transcranial direct current stimulation (tDCS) on food craving, reward and appetite in a healthy population.

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- 1 The effect of transcranial direct current stimulation (tDCS) on food craving, reward
- 2 and appetite in a healthy population.
- 3
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12 ABSTRACT

13 The ability to control hedonic appetite is associated with executive functioning originating in the prefrontal cortex (PFC). The rewarding components of food can override homeostatic 14 15 mechanisms, potentiating obesogenic behaviours. Indeed, those susceptible to 16 overconsumption appear to have PFC hypo-activation. Transcranial direct current 17 stimulation (tDCS) over the dorsolateral PFC (DLPFC) has been shown to reduce food 18 craving and consumption, potentially via attenuating this reward response. We examined the 19 effects of stimulation on food reward and craving using a healthy-weight cohort. This study is 20 amongst the first to explore the effects of tDCS on explicit and implicit components of reward 21 for different food categories. Twenty-one healthy-weight participants (24 ± 7 years, $22.8 \pm$ 22 2.3 kg·m⁻²) completed two sessions involving double-blind, randomised and counterbalanced 23 anodal or sham tDCS over the right DLPFC, at 2 milliampere for 20 minutes. Food craving 24 (Food Craving Questionnaire-State), reward (Leeds Food Preference Questionnaire), and 25 subjective appetite (100 mm visual analogue scales) were measured pre- and post-tDCS. 26 Eating behaviour trait susceptibility was assessed using the Three Factor Eating 27 Questionnaire-Short Form, Control of Eating Questionnaire, and Food Craving 28 Questionnaire-Trait-reduced. Stimulation did not alter food craving, reward or appetite in 29 healthy-weight participants who displayed low susceptibility to overconsumption, with low 30 trait craving, good craving control, and low uncontrolled eating and emotional eating 31 behaviour. Implicit and explicit reward were reliable measures of hedonic appetite, suggesting these are robust targets for future tDCS research. These findings suggest that 32 33 applying tDCS over the DLPFC does not change food reward response in individuals not at 34 risk for overconsumption, and future work should focus on those at risk of overconsumption 35 who may be more responsive to the effects of tDCS on hedonic appetite.

37 KEYWORDS

38 Appetite control; Dorsolateral prefrontal cortex; Neuromodulation; Food craving; Food

39 reward

40

41 HIGHLIGHTS

- We consider the effects of tDCS on implicit and explicit reward using a validated task
- High reliability in reward measures suggest a robust target for future tDCS studies
- Previously findings are limited by high variation within food-related variables
- Effects of tDCS may be dependent on participant eating behaviour traits
- Future work should screen participants using validated psychometric questionnaires

47 1. INTRODUCTION

48 Obesity is a global health epidemic that affects more than 650 million adults worldwide, and 49 is associated with an increased risk of developing many other health conditions (World Health Organisation, 2020). The aetiology of obesity involves a complex relationship 50 51 between behavioural, biological and environmental factors, contributing to the dysregulation 52 of energy balance (Hill, 2006). Hedonic appetite can potentiate this dysregulation, with the 53 rewarding components of food overriding homeostatic mechanisms (Boswell & Kober, 2016; 54 Kober & Boswell, 2018). The ability to control hedonic appetite is associated with executive 55 functioning, which originate in the prefrontal cortex (PFC) and inhibit impulsive actions in 56 favour of goal-directed behaviours (Joseph, Alonso-Alonso, Bond, Pascual-Leone, & 57 Blackburn, 2011). Altered PFC activity in response to food stimuli has been identified in 58 individuals with obesity, especially those displaying binge eating symptoms (Boeka & 59 Lokken, 2011; Karhunen, et al., 2000). It is proposed that a reduction of activity in the right 60 dorsolateral PFC (DLPFC) could facilitate obesogenic behaviours through poor appetite 61 control (Alonso-Alonso & Pascual-Leone, 2007). Indeed, dysregulation of the DLPFC has 62 been linked with greater impulsive behaviours, often leading to overconsumption (Gluck, 63 Viswanath, & Stinson, 2017). Increasing DLPFC activity may improve the ability to control 64 hedonic appetite, providing a novel paradigm in the treatment of obesity (Alonso-Alonso, 65 2013).

66

67 The modulation of cortical activity is possible through the use of non-invasive brain 68 stimulation techniques, such as transcranial direct current stimulation (tDCS). This form of 69 stimulation involves the application of a weak electrical current, typically up to 2 milliampere 70 (mA), to a specific region of the brain via two electrodes that are placed over the scalp 71 (Nitsche & Paulus, 2000). The current is emitted from a battery-powered device, where it is 72 delivered to the brain through an anode electrode and returns to the device through a 73 cathode electrode. The current intensity is not sufficient to cause neuronal firing, but results 74 in the polarity-dependent subthreshold modulation of resting membrane potentials (Filmer,

75 Dux, & Mattingley, 2014; Jamil & Nitsche, 2017). Although the exact mechanisms are not 76 fully understood, it appears the current inhibits neurotransmitters at the synapse; the anode 77 is associated with the inhibition of gamma-aminobutyric acid (GABA) whereas the cathode is 78 associated with the inhibition of glutamate (Filmer, et al., 2014; Stagg, Antal, & Nitsche, 79 2018). The inhibition of these neurotransmitters increases or decreases the likelihood of 80 spontaneous neuronal firing, respectively. In addition to these acute effects, tDCS also 81 appears to elicit changes in cortical activity beyond the stimulation duration. For example, in 82 an early study by Nitsche and Paulus (2001), anodal tDCS lasting 13 minutes resulted in 83 greater activity in the motor cortex for up to 90 minutes post-stimulation.

84

85 When identifying the effects of tDCS on hedonic appetite, many studies have focussed on 86 measuring state food craving. The first study to identify the impact of tDCS on hedonic 87 appetite compared anodal stimulation to the left and right DLPFC in 21 healthy-weight 88 individuals with frequent food cravings, defined as experiencing 3 or more strong urges to 89 consume high-calorie foods per day (Fregni, et al., 2008). When applying 2 mA stimulation 90 for 20 minutes, a significant reduction in food craving was observed following tDCS over the 91 right DLPFC, but not when applied to the left hemisphere. This reduction in state craving 92 score was replicated in a second study that used the same stimulation parameters and 93 recruited a similar participant cohort (n = 19) (Goldman, et al., 2011).

94

95 In two recent publications, Burgess and colleagues highlight a potential eating behaviour 96 trait-dependent effect of tDCS (Burgess, et al., 2016; Ray, et al., 2017). Thirty participants 97 who were obese and met the diagnostic criteria for binge eating disorder (BED) underwent 98 20 minutes of 2 mA tDCS to the right DLPFC, which resulted in a significant decrease in 99 state food craving and in-laboratory food consumption (Burgess, et al., 2016). In contrast, 100 these effects were not significant when this protocol was replicated in 18 participants with 101 frank obesity (i.e. non-binge eating) (Ray, et al., 2017). This suggests that the effects of 102 tDCS may be dependent on individual variation in the level of susceptibility to reward-driven

103 overconsumption. Consistent with this, previous research has demonstrated that individuals 104 with BED are hyper-responsive to the rewarding aspects of food (Davis, 2013; Davis, et al., 105 2009). The estimated prevalence of BED in the general population ranges from 0.7 - 3.0%. 106 and is commonly comorbid with overweight and obesity (Kessler, et al., 2013). Recurrent 107 episodes of binge eating behaviour are estimated to occur in 10 - 20% of individuals who 108 are healthy weight, overweight or obese, and constitutes a trait that can be assessed 109 psychometrically and applied to a non-clinical population. Similar to findings in individuals 110 with BED, individuals with eating behaviour trait susceptibility to overconsume (i.e. binge 111 eating and emotional eating) have been found to be hyper-responsive to the rewarding 112 aspects of food (Dalton, Blundell, & Finlayson, 2013a). Therefore, including validated 113 measures of food reward and eating behaviour trait susceptibility may be important when 114 considering the effect of tDCS on food consumption, reward and craving. To date, no study 115 has identified the effects of tDCS on implicit and explicit components of reward across 116 different food categories.

117

118 Although there are many promising findings, not all studies have found an effect of tDCS on 119 measures of hedonic appetite. This may be due to the inconsistent application of stimulation 120 parameters (e.g. variation in target electrode placement and current intensity), inadequate 121 experimental blinding, and large variation in experimental measures (Hall, Vincent, & 122 Burhan, 2018; Tremblay, et al., 2014). The most consistently used measure of hedonic 123 appetite in tDCS research is food craving, which is commonly assessed using the Food 124 Craving Questionnaire-State (FCQ-S) (Cepeda-Benito, Gleaves, Williams, & Erath, 2000). 125 Although significant effects of tDCS on food craving have been identified (Fregni, et al., 126 2008; Goldman, et al., 2011), this has not been consistently shown (Georgii, Goldhofer, 127 Meule, Richard, & Blechert, 2017; Sedgmond, et al., 2019). Across studies there is large 128 variation in state food craving scores, ranging from 0.40% to 41.67% following the active 129 condition (Fregni, et al., 2008; Goldman, et al., 2011; Kekic, et al., 2014; Ljubisavljevic, 130 Maxood, Bjekic, Oommen, & Nagelkerke, 2016), which may be due to the poor reliability of

131 these measures. Developmental publications of the FCQ-S suggest low-to-moderate 132 reliability (r = 0.39 - 0.56) (Cepeda-Benito, et al., 2000; Meule, Teran, et al., 2014). 133 Measures of food consumption have also been utilised, primarily using ad libitum buffets of 134 highly palatable foods (Burgess, et al., 2016; Georgii, et al., 2017; Gluck, et al., 2015; Ray, 135 et al., 2017; Sedgmond, et al., 2019). Although greater craving control is associated with 136 improved weight loss outcomes (Dalton, et al., 2017), the effects of tDCS on craving and consumption are not correlated (Burgess, et al., 2016), suggesting other targets are required 137 138 to validate tDCS as an intervention to alter eating behaviour. Food reward plays a more 139 pivotal role in the dysregulation of energy balance (Boswell & Kober, 2016; Kober & Boswell, 140 2018). Therefore, it is important to look beyond the measure of food craving and identify the 141 role of tDCS in modulating food reward.

142

143 The present study examined how measures of food craving, reward and appetite would 144 change following the inducement of hyper-activity of the right DLPFC through tDCS in a 145 healthy-weight cohort. We hypothesised stimulation would reduce state food craving and 146 subjective appetite, based on previous findings utilising healthy participant groups (Fregni, et 147 al., 2008; Goldman, et al., 2011; Kekic, et al., 2014; Lapenta, Sierve, de Macedo, Fregni, & 148 Boggio, 2014). We also hypothesised that participants' preference for and perceived 149 rewarding value of high-fat and sweet foods would be diminished following anodal tDCS. We 150 also looked to establish the reliability of these measures, including both implicit and explicit 151 components of reward, prior to tDCS with a view to establishing the viability of their future 152 use in our research.

153

154 **2. METHODS**

155 2.1 Participants

156 The study was approved by an institutional ethics committee, and all participants provided

157 written informed consent. Sample size was determined using G*Power 3.0.10 (Faul,

158 Erdfelder, Lang, & Buchner, 2007). An effect size f of 0.33 was based on mean percentage

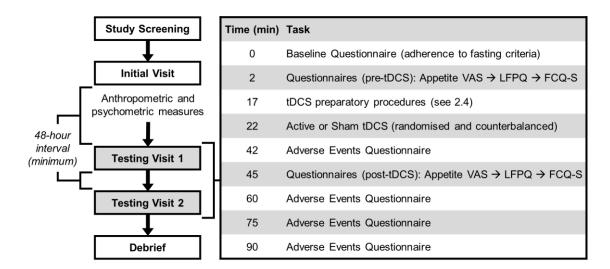
159 difference from baseline in food craving scores following single session tDCS (mean 160 difference between conditions -22.22 ± 33.68%) (Fregni, et al., 2008; Goldman, et al., 2011; 161 Kekic, et al., 2014; Ljubisavljevic, et al., 2016). Using α error probability of 0.05, 1 group with 162 2 measurements, a correlation among repeated measures equal to 0.5, and non-sphericity 163 correlation E of 1, sample size calculations determined a minimum sample size of 21, with 164 actual power of 0.82, was required. Twenty-one participants $(24 \pm 7 \text{ years}, 22.8 \pm 2.3 \text{ kg} \cdot \text{m}^{-2})$ 165 were recruited via email and poster advertisements. Interested individuals were initially 166 screened with an online questionnaire. Eligible participants were male or female between 18 167 and 60 years of age who presented with a body mass index (BMI) between 18.5 and 24.9. 168 All participants were free of neurological, cardiovascular, metabolic and joint disease, and 169 potential participants were excluded if they presented with low mood or depressive 170 symptoms, as indicated using the Centre for Epidemiologic Studies Short Depression Scale 171 (CESD-10) (Andresen, Malmgren, Carter, & Patrick, 1994; Radloff, 1977). Female 172 participants who were pregnant or wishing to conceive were also excluded from the study. 173 Participants were naïve to tDCS protocols, non-smokers and were not recreational drug 174 users or taking any medications at the time of data collection.

175

176 2.2 Experimental Design

177 The study utilised a double-blind, within-participant, repeated measures design. Participants 178 attended the laboratory on three separate occasions (Figure 1). During the first visit, 179 participants completed a series of psychometric questionnaires, and height and body 180 composition were measured. Visits 2 and 3 were experimental trials where all participants 181 received either active or sham tDCS in a randomised and counterbalanced order. 182 Randomisation was determined using a permuted block paradigm and completed by an 183 independent party. The participants and researcher conducting stimulation were blind to the 184 tDCS condition, adhering to a double-blind design, which was maintained using a pin-185 protected device. All sessions were scheduled at the same time of day within-subject, 186 occurring between 09:00 and 15:00, and with a minimum interval of 48 hours between

- 187 sessions to prevent any residual effects of stimulation (Alonzo, Brassil, Taylor, Martin, & Loo,
- 188 2012).
- 189
- 190



192 Figure 1 Study Procedure

193 FCQ-S, Food Craving Questionnaire-State; LFPQ, Leeds Food Preference Questionnaire;

194 tDCS, transcranial direct current stimulation; VAS, visual analogue scales.

195

196 2.3 Procedure

Participants were required to fast for a minimum of 4 hours prior to each visit, where they were asked to refrain from consuming any food or drink other than water. In addition, they were asked to refrain from consuming products containing caffeine and alcohol in the 12 or 24 hours prior to each visit, respectively. Adherence to this fasting criteria was self-reported at the start of each visit. On arrival at the laboratory, participants were instructed to turn off their mobile phones and remove any metallic objects from their person in adherence with our tDCS protocol.

- 205 During visit 1, participants completed the Three Factor Eating Questionnaire-Short Form
- 206 (TFEQ-r18), Control of Eating Questionnaire (CoEQ), and Food Craving Questionnaire-Trait-
- 207 reduced (FCQ-T-r); see 2.5.1. Height was measured using a portable stadiometer (SECA

208 Limited, Birmingham, UK) to the nearest mm. Measurements were taken following 209 inhalation, with the participant standing straight, their back to the rule, and their eyeline level 210 with ear canal. Body composition, including weight and BMI, was assessed using a Tanita 211 BC-418MA analyser (Tanita Europe B.V., Amsterdam). Weight was measured to the nearest 212 0.1 kg, and body fat percentage to the nearest 0.1%. Participants were then shown the food images used in the Leeds Food Preference Questionnaire (LFPQ; see 2.5.4), and their 213 214 familiarity and acceptance of each food item was assessed. Any food items that were 215 unfamiliar or had low acceptance (i.e. disliked or not consumed in the normal diet) were 216 substituted with images from a database of additional items with similar nutritional and 217 sensory properties (Oustric, et al., 2020).

218

219 During visits 2 and 3, participants completed a series of questionnaires immediately pre-220 tDCS (Figure 1). These included appetite visual analogue scales (VAS), the FCQ-S and the 221 LFPQ; see 2.5. Each participant then underwent 20 minutes of active or sham tDCS. 222 Questionnaires were then repeated immediately post-stimulation. Visits 2 and 3 were 223 identical, apart from the stimulation condition. At the end of visit 3, participants were 224 debriefed. They were informed of the sham stimulation condition and were asked whether 225 they were able to differentiate between the active and sham conditions, and in which visit 226 they believe active tDCS was delivered.

227

228 2.4 Stimulation Protocol

229 Stimulation was delivered using the HDCstim direct current stimulator (Newronika s.r.l.,

230 Milan, Italy) by a trained researcher. Anodal stimulation was used to target the right DLPFC,

- in accordance with the International Standards for Electroencephalography 10-20 system
- 232 (Klem, Lüders, Jasper, & Elger, 1999). A 25 cm² anode electrode was placed over the frontal
- area 4 (F4) and a 51 cm² cathode electrode placed over occipital zero point (Oz). Cathode
- 234 placement over the Oz reduces the impact of associated inhibitory effects on study
- measures (Bestmann, de Berker, & Bonaiuto, 2015; Galetta, 2017), and decreases the

236 likelihood of current shunting across the scalp (Rush & Driscoll, 1968). Rubber electrodes 237 were housed in sponge pads, pre-soaked in 0.9% sodium chloride. A constant current of 2 238 mA was delivered through the anode electrode, culminating in a current density of 0.08 239 mA·cm⁻². The current was ramped up over a 30-second period, and active tDCS was then 240 delivered for 20 minutes, with a 30-second ramp down. Stimulation was delivered offline (i.e. 241 no task was performed during tDCS), and participants were instructed to remain seated. 242 relaxed and awake. Sham stimulation involved the same setup as active tDCS, but the 243 current was only delivered for 36 seconds (3% active tDCS duration). This is associated with 244 similar sensations (e.g. itching, tingling) (Brunoni, et al., 2011; Nikolin, Huggins, Martin, 245 Alonzo, & Loo, 2018), but has a limited neuromodulatory effect (Gandiga, Hummel, & 246 Cohen, 2006).

247

The effectiveness of sham as a blinding technique was assessed during debrief. Impedance
was measured at the start of stimulation, and periodically checked thereafter. It is
recommended that impedance should remain below 5 kilo-ohm (kΩ) (DaSilva, Volz, Bikson,
& Fregni, 2011; Thair, Holloway, Newport, & Smith, 2017). The occurrence of sensations
and adverse events were measured using a standardised questionnaire (Brunoni, et al.,
2011) immediately following stimulation and at regular intervals for a minimum of 45 minutes,
in accordance with our standardised procedure.

255

256 2.5 Measurements

257 2.5.1 Psychometric Questionnaires

A series of psychometric questionnaires were used to assess eating behaviour traits of

259 participants in the screening session (Figure 1). The TFEQ-r18 (Karlsson, Persson,

- Sjöström, & Sullivan, 2000) measures three subscales of eating behaviour; cognitive
- restraint, uncontrolled eating, and emotional eating. Scores range from 0 to 100 for each
- subscale, with higher scores indicating a greater presence of problematic eating behaviour.
- 263 The TFEQ-r18 has good internal validity, with a Cronbach's alpha (α) of 0.82, and

comparable construct validity to the full TFEQ (r = 0.71 - 0.99) (Karlsson, et al., 2000). The 264 265 CoEQ (Dalton, Finlayson, Hill, & Blundell, 2015) considers the frequency, intensity and 266 severity of food cravings experienced over the previous 7 days. Items are assessed using 267 100 mm VAS, with scores averaged across items to provide an individual score for craving 268 control, craving for sweet foods, craving for savoury foods, and positive mood. Cronbach's a 269 highlighted acceptable internal validity; craving control $\alpha = 0.88$, craving for sweet foods $\alpha =$ 270 0.67, craving for savoury foods α = 0.66, positive mood α = 0.74. Finally, general and 271 habitual food cravings were measured using the 15-item FCQ-T-r (Meule, Hermann, & 272 Kübler, 2014). This questionnaire assesses lack of control over eating, emotions 273 experienced before or during food craving and consumption, and guilt from cravings and/or 274 giving in to cravings. A higher score suggests more frequent cravings and a total score 275 greater than 50 highlights clinically relevant trait cravings (Meule, 2018). The FCQ-T-r has 276 high internal validity ($\alpha = 0.94$).

277

278 2.5.2 Appetite Visual Analogue Scales (VAS)

Four 100 mm VAS were used to assess subjective ratings of appetite (Blundell, et al., 2010),
which are sensitive to experimental manipulation and considered reliable and valid
measures of subjective appetite (Beechy, Galpern, Petrone, & Das, 2012). Scales measured
hunger (*"How hungry do you feel right now?"*), fullness (*"How full do you feel right now?"*),
prospective consumption (*"How much food could you eat right now?"*), and the desire to eat
(*"How strong is your desire to eat right now?"*). Scores range from 0 to 100, with higher
scores indicating greater prevalence of the appetite measure.

286

287 2.5.3 Food Craving Questionnaire-State (FCQ-S)

The FCQ-S (Cepeda-Benito, et al., 2000) measures subjective food craving, and is
responsive to situational changes (Cepeda-Benito, et al., 2000; Meule, Teran, et al., 2014).
This questionnaire assesses the desire to eat, craving for food, and emotional responses to

291 food and consumption over 15 statements. Participants rate each statement on a 5-point

scale, where 1 corresponds with "*Strongly disagree*" and 5 corresponds with "*Strongly agree*". Corresponding scores are totalled, with a minimum score of 15 and a maximum of 75; higher scores equating to greater momentary craving. Similar to the FCQ-T-r, the state FCQ has good internal validity (Cronbach's $\alpha = 0.94$) (Cepeda-Benito, et al., 2000).

296

297 2.5.4 Leeds Food Preference Questionnaire (LFPQ)

298 The LFPQ (Dalton & Finlayson, 2014; Finlayson, King, & Blundell, 2007) is a validated 299 computer-based assessment of hedonic preference for food, measuring explicit liking and 300 wanting and implicit wanting as components of reward. "Liking" can be defined as the 301 subjective pleasure elicited by food or related cues, whereas "wanting" is the motivational 302 component of reward that refers to subjective desire or craving for foods (see Finlayson and 303 Dalton (2012) for review). Liking operates at an explicit level (i.e. conscious, introspective), 304 and wanting at both explicit and implicit (i.e. subconscious, automatic) levels (Finlayson & 305 Dalton, 2012). The task uses a standardised set of 16 images depicting ready-to-eat foods 306 that are common in the diet (Table 1), and reward is assessed according to the fat content 307 and taste of these foods. Food images illustrate items that are either high (>40% energy) or 308 low (<20% energy) in fat, and either sweet or savoury. Food items are split into four 309 categories; high-fat savoury (HFSA), high-fat sweet (HFSW), low-fat savoury (LFSA), and 310 low-fat sweet (LFSW). The food items are comparable in protein content, palatability and 311 familiarity (Oustric, et al., 2020).

313	Table 1	Standardised f	ood images	used in the LFPQ	Į.
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HFSA	HFSW	LFSA	LFSW
Garlic bread	Chocolate biscuits	Green salad	Mixed berry salad
Fries	Glazed doughnut	Broccoli	Skittles
Crisps	Blueberry muffin	Vegetable rice	Haribo
Sausage roll	Milk chocolate	Bread roll	Banana

HFSA, high-fat savoury; HFSW, high-fat sweet; LFSA, low-fat savoury; LFSW, low-fat sweet

314

315 The LFPQ involves two tasks, where food items are displayed in pairs (forced-choice task) 316 or individually (single-food task). The forced-choice task measures the implicit wanting for 317 foods. Participants are required to choose the food they most want to consume "right now" 318 from two items presented on a computer screen. Scores for implicit wanting are calculated 319 using a frequency-weighted algorithm, by combining reaction time and the frequency of 320 choosing or avoiding a food (Dalton & Finlayson, 2014). In the single-food task, participants 321 are presented with each of the 16 food items individually and asked to rate "How much do 322 you want some of this food right now?" and "How pleasant would it be to taste some of this 323 food right now?" on 100 mm VAS. This second task measures explicit wanting and liking, 324 respectively, for each food item. In addition, fat appeal bias (FAB) and taste appeal bias 325 (TAB) scores are calculated by subtracting mean scores across food groups (e.g. mean low-326 fat scores subtracted from mean high-fat scores), and provide scores for explicit liking, 327 explicit wanting and implicit wanting. 328 329 2.6 Data Analysis

330 Mean and standard deviations (SD) were calculated at each time point (pre- and post-

331 stimulation) under active and sham tDCS conditions. Normality of data were assessed using

332 Shapiro-Wilks test, and reliability of baseline measures were determined using Pearson's *r*

correlations. The effects of tDCS on FCQ-S and LFPQ scores were evaluated using a 2 333 334 (condition; active or sham) * 2 (time point; pre-stimulation vs post-stimulation scores) 335 repeated-measures analysis of variance (ANOVA). Post-hoc significant effects were 336 determined using pair-wise comparisons with Bonferroni correction. Although fasting 337 protocols were standardised, significant differences in baseline scores across all appetite 338 VAS measures were found. To control for this difference, scores were transformed to 339 difference from baseline and analysed using a paired-samples t-test. To determine the 340 impact of difference in baseline hunger scores on other test variables, analysis of covariance 341 (ANCOVA) were performed with baseline hunger as a covariate. ANCOVA were additionally 342 performed to control for behaviour trait scores. Adverse events were compared using further 343 paired-samples t-tests. Statistical analyses were performed using SPSS version 21 and 26 344 (IBM, New York, USA). Data are presented as mean \pm SD to an alpha level of 0.05.

345

To interpret the null findings and assess the strength of evidence, Bayesian statistics were

347 computed using JASP (version 0.13.1; University of Amsterdam). The classification scheme

by Lee and Wagenmakers (2013) provides descriptive labels for Bayes factors (BF₁₀), and

349 was used to interpret values. In brief, scores greater than 1 provide evidence in favour of the

350 alternative hypothesis, whereas scores below 1 provide evidence in favour of the null

351 hypothesis. Scores are labelled as anecdotal (score between 1 and 3 or 1 and 0.33),

moderate (score between 3 and 10 or 0.33 and 0.10), strong (score between 10 and 30 or

- 0.10 and 0.03), very strong (score between 30 and 100 or 0.03 and 0.01), or extreme (score
- 354 greater than 100 or lesser than 0.01).

355

356 3. RESULTS

Participant anthropometric and psychometric characteristics are displayed in
 Table 2. Participants were weight stable (±5%) for 3 months prior to the study, and displayed

359 "healthy" eating behaviour trait profiles as identified by FCQ-T, CoEQ and TFEQ-r18 scores.

360 Scores for the FCQ-T-r were below the 50-point cut-off for clinically relevant trait craving

- 361 (Meule, 2018), with CoEQ and TFEQ-r18 scores comparable to healthy-weight individuals in
- 362 other studies (Anglé, et al., 2009; De Lauzon-Guillain, et al., 2009; Fleurbaix Laventie Ville
- 363 Sante Study, 2004; Wardle, et al., 2018).

	Female	Male	All
Ν	11	10	21
Age (years)	25 ± 9 (19 – 52)	23 ± 4 (20 – 29)	24 ± 7 (19 – 52)
Height (cm)	165 ± 6 (155 – 175)	179 ± 6 (170 – 189)	172 ± 9 (155 – 189)
Weight (kg)	60.1 ± 7.4 (49.6 – 71.4)	76.5 ± 7.1 (66.6 – 88.9)	67.9 ± 11.0 (49.6 – 88.9)
BMI (kg⋅m ⁻²)	22.0 ± 2.1 (18.5 – 25.0)	23.8 ± 2.2 (20.1 – 27.7*)	22.8 ± 2.3 (18.5 – 27.7*)
Body fat (kg)	16.3 ± 4.3 (10.9 – 23.3)	12.9 ± 4.9 (6.4 – 20.7)	14.7 ± 4.8 (6.4 – 23.3)
Body fat (%)	26.8 ± 4.3 (20.6 – 33.1)	16.6 ± 5.5 (9.2 – 26.0)	21.9 ± 7.1 (9.2 – 33.1)
CESD-10 (AR)	5 ± 3 (0 – 10)	5 ± 4 (0 – 10)	5 ± 3 (0 – 10)
FCQ-T-r (AR)	36 ± 8 (22 – 49)	34 ± 10 (20 – 47)	35 ± 9 (20 – 49)
TFEQ-r18 Cognitive Restraint (AR)	34 ± 19 (5.6 – 61.1)	33 ± 21 (11.1 – 77.8)	40 ± 20 (5.6 – 77.8)
TFEQ-r18 Uncontrolled Eating (AR)	33 ± 11 (7.4 – 44.4)	34 ± 18 (3.7 – 66.7)	33 ± 14 (3.7 – 66.7)
TFEQ-r18 Emotional Eating (AR)	24 ± 24 (0.0 - 66.7)	20 ± 23 (0.0 - 66.7)	22 ± 22 (0 – 66.7)
CoEQ Craving Control (mm)	66 ± 18 (36.0 – 96.2)	68 ± 18 (36.4 – 94.1)	65 ± 18 (36.0 – 96.2)
CoEQ Craving for Sweet Foods (mm)	30 ± 16 (3.0 – 59.7)	28 ± 21 (2.3 – 67.0)	29 ± 18 (2.3 – 67.0)
CoEQ Craving for Savoury Foods (mm)	54 ± 19 (16 – 78)	46 ± 26 (2.0 – 79.3)	51 ± 23 (2.0 – 79.3)
CoEQ Positive Mood (mm)	51 ± 16 (20 – 84)	54 ± 13 (31.0 – 68.3)	52 ± 14 (20 – 84)

365 **Table 2** Mean, standard deviation and range for participant anthropometric and eating behaviour trait characteristics

* n = 1 with BMI >24.9 due to high fat-free mass (weight 88.9 kg, fat-free mass 74.2 kg, fat mass 14.7 kg / 16.5%).

BMI, Body Mass Index; CESD-10, Centre for Epidemiologic Studies Short Depression Scale; FCQ-T-r, Food Craving Questionnaire-Trait reduced form; TFEQ-r18, Three Factor Eating Questionnaire 18-item version; CoEQ, Control of Eating Questionnaire.

368 3.1 Appetite Visual Analogue Scales (VAS) 369 Baseline hunger scores were significantly higher in the active session $(63.1 \pm 21.4 \text{ mm})$, 370 when compared to the sham session (51.9 \pm 25.8 mm) (t₍₂₀₎ = 2.567, p = 0.018). Similarly, 371 scores for fullness ($t_{(20)} = 2.925$, p = 0.008), prospective consumption ($t_{(20)} = 3.196$, p = 372 0.005), and desire to eat ($t_{(20)} = 2.756$, p = 0.012) were greater at baseline in the active 373 versus sham session. Baseline hunger scores did not significantly affect fullness, 374 prospective consumption or desire to eat (p's > 0.05). There were no significant changes in subjective hunger ($t_{(20)} = 0.572$, p = 0.574, BF₁₀ = 0.264), fullness ($t_{(20)} = 0.146$, p = 0.886, 375 $BF_{10} = 0.230$), prospective consumption (t₍₂₀₎ = 0.969, p = 0.344, BF_{10} = 0.345), or desire to 376 377 eat ($t_{(20)} = 1.772$, p = 0.092, BF₁₀ = 0.858) when comparing pre- and post-stimulation in the 378 active and sham tDCS conditions (Figure 2). Bayes factors show moderate evidence in 379 favour of the null hypothesis over the alternative hypothesis for hunger and fullness, 380 whereas prospective consumption and the desire to eat were supported only by anecdotal 381 evidence in favour of the null hypothesis. When controlling for behaviour traits scores, the 382 effects of tDCS on the desire to eat only neared significance (p = 0.062 - 0.076), and 383 remained non-significant for other subjective appetite measures (p's > 0.32).

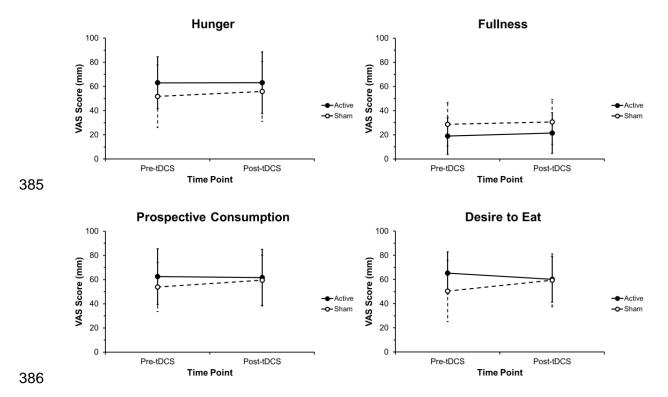


Figure 2 Mean \pm SD appetite visual analogue scale (VAS) scores prior to and following tDCS intervention (n = 21).

390 **3.2 Food Craving Questionnaire-State (FCQ-S)**

There were no significant differences in state food craving from pre- to post-stimulation under active (pre-tDCS 47.2 ± 9.9 AU, post-tDCS 47.8 ± 12.2 AU) or sham conditions (pretDCS 43.8 ± 10.2 AU, post-tDCS 44.9 ± 9.0 AU) ($F_{(1, 19)} = 0.069$, p = 0.797). Bayes factor highlights moderate evidence in favour of the null hypothesis over the alternative hypothesis (BF₁₀ = 0.272). In addition, this effect remained non-significant when controlling for baseline hunger ($F_{(1, 38)} = 0.037$, p = 0.849) and behaviour trait scores (p > 0.74).

397

398 **3.3 Leeds Food Preference Questionnaire (LFPQ)**

Stimulation did not alter measures of implicit or explicit food reward, with the condition * time point interactions for the liking and wanting of HFSA, LFSA, HFSW and LFSW categories showing no significant effect (p > 0.05) (Table 3), which is supported by moderate-to-strong evidence in favour of the null hypothesis over the alternative hypothesis (BF₁₀ = 0.041 –

403 0.168). The interactions remained non-significant when controlling for baseline hunger (p's > 404 0.10) and behaviour trait scores (p's > 0.11). In addition, tDCS did not significantly change implicit or explicit TAB, with non-significant condition * time point interactions for explicit 405 406 liking ($F_{(1, 18)} = 0.079 \text{ p} = 0.782$, $BF_{10} = 0.030$), explicit wanting ($F_{(1, 18)} = 0.902$, p = 0.355, 407 $BF_{10} = 0.078$), and implicit wanting ($F_{(1, 17)} = 0.786$, p = 0.388, $BF_{10} = 0.076$). Again, this is supported by strong evidence in favour of the null hypothesis over the alternative hypothesis 408 409 and the effects remained non-significant when controlling for baseline hunger (p's > 0.40) 410 and behaviour trait scores (p's > 0.42). Similar non-significant condition * time point 411 interactions were seen for FAB explicit wanting ($F_{(1, 18)} = 0.136$, p = 0.716, $BF_{10} = 0.183$) and 412 implicit wanting ($F_{(1, 17)} = 0.646$, p = 0.433, BF₁₀ = 0.111). These scores remained non-413 significant when controlling for baseline hunger (p = 0.823 and 0.236, respectively) and 414 behaviour trait scores (p's > 0.24). However, there was a significant time point ($F_{(1, 18)}$ = 6.785, p = 0.018) and condition * time point interaction for FAB explicit liking ($F_{(1, 18)} = 7.374$, 415 416 p = 0.014, $BF_{10} = 0.545$); scores increased following both active and sham tDCS, and to a 417 greater extent following active stimulation (Table 3). After controlling for baseline hunger 418 scores this effect was no longer significant ($F_{(1, 36)} = 2.944$, p = 0.095, $BF_{10} = 0.313$). 419 Similarly, when controlling for baseline behaviour trait scores no significant effects were 420 identified (p's > 0.08).

		Explicit Liking	(mm)	Explicit Wanti	ng (mm)	Implicit Wantir	ng (AU)
		Pre-tDCS	Post-tDCS	Pre-tDCS	Post-tDCS	Pre-tDCS	Post-tDCS
Active	HFSA	52.9 ± 23.9	64.7 ± 25.1	49.9 ± 23.8	59.9 ± 25.7	-0.7 ± 31.0	15.7 ± 47.7
	LFSA	54.0 ± 20.9	53.7 ± 22.0	53.3 ± 22.1	53.6 ± 19.3	-1.4 ± 24.9	-15.1 ± 38.7
	HFSW	48.2 ± 24.6	54.6 ± 23.0	48.0 ± 24.2	48.8 ± 24.4	-6.5 ± 27.9	-3.6 ± 26.6
	LFSW	60.2 ± 20.8	60.5 ± 21.8	60.0 ± 19.5	57.1 ± 19.5	10.8 ± 28.7	3.8 ± 26.5
	FAB	-6.6 ± 24.1*	2.5 ± 21.4*	-7.8 ± 25.0	-1.0 ± 23.3	-7.1 ± 45.3	12.1 ± 53.7
	TAB	0.7 ± 18.2	-1.6 ± 20.0	2.4 ± 15.5	-3.9 ± 19.3	3.8 ± 20.2	-9.5 ± 39.1
Sham	HFSA	53.8 ± 26.6	57.2 ± 25.2	51.8 ± 27.7	55.1 ± 25.8	9.6 ± 33.1	14.4 ± 28.5
	LFSA	49.7 ± 18.4	49.8 ± 18.1	49.7 ± 20.1	49.8 ± 18.2	-2.3 ± 25.2	-3.6 ± 23.5
	HFSW	49.0 ± 27.6	47.1 ± 26.4	42.6 ± 28.0	45.9 ± 26.1	-9.5 ± 29.5	-6.9 ± 30.9
	LFSW	57.4 ± 22.3	55.4 ± 19.0	56.8 ± 20.1	53.5 ± 20.5	5.5 ± 30.0	-0.6 ± 29.4
	FAB	-2.1 ± 26.3*	-0.4 ± 23.2*	-6.1 ± 29.6	-1.1 ± 23.8	0.2 ± 45.0	7.6 ± 41.9
	TAB	1.4 ± 18.9	-2.3 ± 12.9	-1.0 ± 16.3	-2.8 ± 11.6	-5.3 ± 29.4	-8.7 ± 25.6

422	Table 3 Pre-stimulation and post-stimulation LFI	PQ scores ($n = 21$).
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Mean ± SD. HFSA, high-fat savoury; LFSA, low-fat savoury; HFSW, high-fat sweet; LFSW, low-fat sweet; FAB, fat appeal bias; TAB,

taste appeal bias. * Indicates significant difference between active and sham conditions (p < 0.05).

425 3.4 Test-Retest Analysis

- 426 With the exception of desire to eat (r = 0.382, p = 0.088), all variables were significantly
- 427 correlated between baseline assessment. Twelve of the 23 variables assessed (across
- 428 measurement instruments) displayed a strong correlation (r = >0.7), with LFPQ implicit
- 429 wanting and FAB appearing most consistent. Some baseline measures, particularly FCQ-S
- 430 and appetite VAS measures, indicated only moderate reliability (r = 0.5 to 0.7; Table 4).

			r	р
Food Craving	Questionnaire-State		0.549	0.010
Appetite VAS	Hunger		0.654	0.001
	Fullness		0.588	0.005
	Prospective Consumption		0.841	<0.001
	Desire to Eat		0.382	0.088
LFPQ	Implicit Wanting	HFSA	0.837	<0.001
		LFSA	0.795	<0.001
		HFSW	0.882	<0.001
		LFSW	0.718	0.001
	Explicit Liking	HFSA	0.652	0.002
		LFSA	0.664	0.002
		HFSW	0.781	<0.001
		LFSW	0.784	<0.001
	Explicit Wanting	HFSA	0.698	0.001
		LFSA	0.751	<0.001
		HFSW	0.712	0.001
		LFSW	0.668	0.002
	Fat Appeal Bias	Explicit Liking	0.853	<0.001
		Explicit Wanting	0.887	<0.001
		Implicit Wanting	0.677	0.001
	Taste Appeal Bias	Explicit Liking	0.536	0.018
		Explicit Wanting	0.555	0.014
		Implicit Wanting	0.737	<0.001

432	Table 4 Correlations between baseline (nra-tDCS) maasuras	during visits 2 and 3 $(n - 21)$
432	able 4 Correlations between baseline (pie-iDCS) measures i	uuning visits z and 5 (n = z r).

HFSA, high-fat savoury; HFSW, high-fat sweet; LFPQ, Leeds Food Preference

Questionnaire; LFSA, low-fat savoury; LFSW, low-fat sweet; VAS, visual analogue scale.

434 3.5 Responses to tDCS

435 Successful delivery of the electric current occurred across all 42 stimulation sessions, with 436 mean impedance levels of 8 ± 4 k Ω at the start of stimulation which reduced to 3 ± 2 k Ω 437 within the initial five minutes of current delivery. Stimulation was well-tolerated by 438 participants with only common adverse events, particularly cutaneous sensations, 439 experienced during tDCS. The most common sensations reported were tingling, itching, 440 sleepiness and a burning sensation (Table 5). Tingling (p = 0.016), itching (p = 0.021) and 441 sleepiness (p = 0.021) were reported by significantly more participants following active 442 tDCS, when compared with sham stimulation. No other sensations were significantly 443 different between conditions. Despite experiencing more minor adverse events, participants 444 were unable to identify the active tDCS session above the level of chance, with only 38% 445 (8/21) of participants able to successfully distinguish between conditions.

	Active	Sham	р
Headache	7 (33%)	4 (19%)	0.186
Neck pain	0 (0%)	0 (0%)	-
Scalp pain	3 (14%)	1 (5%)	0.329
Tingling	14 (67%)	7 (33%)	0.016*
Itching	11 (52%)	6 (29%)	0.021*
Burning sensation	9 (43%)	2 (10%)	0.267
Skin redness	5 (24%)	2 (10%)	0.186
Sleepiness	12 (57%)	7 (33%)	0.021*
Trouble concentrating	5 (24%)	3 (14%)	0.329
Acute mood change	2 (10%)	2 (10%)	1.000

447 **Table 5** Frequency of adverse events experienced immediately post-stimulation.

* Indicates significant difference between active and sham conditions.

448

449 **4. DISCUSSION**

450 The current study examined the effect of induced hyper-activity of the right DLPFC through 451 tDCS on food craving, reward and subjective appetite measures in a healthy-weight cohort. 452 It is important to note that the sample used in the current study appeared to show low 453 susceptibility to hedonic-driven overconsumption, evidenced by their scores on several 454 measures of eating behaviour traits linked to overconsumption. We also sought to examine the reliability of measures prior to tDCS with a view to establishing the viability of their future 455 456 use in our research. We report strong relationships between key variables, particularly implicit wanting and FAB, when preparatory procedures prior to tDCS were standardised. 457 These variables may prove to be sensitive targets for detecting significant effects in future 458 459 eating behaviour-focussed tDCS research. Other variables, particularly food craving 460 measures, proved less stable and may require tighter experimental control or larger sample 461 sizes to reveal differences. Collectively our findings are novel to tDCS research.

463 Prior work has mainly focussed on measuring food craving and in-laboratory consumption 464 with equivocal findings (Fregni, et al., 2008; Georgii, et al., 2017; Goldman, et al., 2011; 465 Sedgmond, et al., 2019). The present study is favourable by comparison in sample size, 466 study design (i.e. sham-controlled and double-blind) and stimulation parameters (Burgess, et 467 al., 2016; Fregni, et al., 2008; Goldman, et al., 2011; Ray, et al., 2017). Recently published 468 meta-analyses have cast doubt in the ability of tDCS to alter measures of food craving (Hall 469 & Lowe, 2018; Lowe, Vincent, & Hall, 2017), which may be due to the poor test-retest 470 reliability of food craving measures (Cepeda-Benito, et al., 2000; Meule, Teran, et al., 2014). 471 This is in agreement with our data which highlighted only moderate reliability of baseline 472 FCQ-S scores. In comparison, our data show strong relationships between baseline 473 measures of implicit and explicit reward. In developing the LFPQ, Dalton and Finlayson 474 (2014) reported a reliability coefficient of 0.6 - 0.7 for implicit wanting and 0.8 - 0.9 for 475 explicit liking measures, with our data supporting this moderate-to-strong reliability. The 476 LFPQ has been utilised in several settings, and is considered a sensitive measure for 477 individual eating behaviour traits (Dalton, et al., 2013a; Dalton, Blundell, & Finlayson, 2013b; 478 Finlayson, Arlotti, Dalton, King, & Blundell, 2011), and a good predictor of in-laboratory and 479 real-world food choice and consumption (French, Mitchell, Finlayson, Blundell, & Jeffery, 480 2014; Griffioen-Roose, Mars, Finlayson, Blundell, & de Graaf, 2011). The present study is 481 the first to extend the use of the LFPQ to include tDCS procedures, and the reliability of this 482 questionnaire suggests it is a robust measure and should be explored in future research.

483

It is logical that the significant interaction between tDCS condition and time point for FAB explicit liking was removed when controlling for baseline hunger as the excitatory effects of anodal tDCS are not associated with increased preference for high-fat foods (Goldman, et al., 2011; Jauch-Chara, et al., 2014). In addition, healthy individuals are likely to have a normative response to food stimuli and are able to sufficiently integrate rewarding signals with other appetitive signals to select appropriate eating behaviours (see Alonso-Alonso and

490 Pascual-Leone (2007) for review). Healthy individuals are also unlikely to have structural 491 deficits observed in obesity and binge eating, which are associated with alteration in reward response (Balodis, Grilo, & Potenza, 2015; Lowe, Reichelt, & Hall, 2019). It is probable that 492 493 stimulation would have no additive effects in healthy individuals (Burgess, et al., 2016). The 494 greater baseline hunger score likely heightened the rewarding value of high-calorie foods, 495 particularly those high in fat, that participants were exposed to during the computer-based 496 task (Cameron, Goldfield, Finlayson, Blundell, & Doucet, 2014; Finlayson, King, & Blundell, 497 2008; Mehta, et al., 2012).

498

499 In addition to the equivocal findings for food craving and consumption, previous work has 500 been inconsistent in the recruitment of participants and some of the variation in results may 501 be due to participants' eating behaviour traits. Two previous studies have directly linked 502 tDCS effects as occurring in those with abnormal eating behaviours (Burgess, et al., 2016; 503 Ray, et al., 2017), and when comparing further studies that utilise similar tDCS parameters 504 (i.e. 2 mA for 20 minutes over the right DLPFC), a trait-dependent effect is evident. Studies 505 that recruited participants with frequent food cravings found a consistent reduction in 506 measures of state food craving (Fregni, et al., 2008; Goldman, et al., 2011; Kekic, et al., 507 2014; Lapenta, et al., 2014). In comparison, studies that did not measure behaviour traits or 508 report comparable traits between healthy and overweight populations, fail to find a significant 509 effect of stimulation on craving (Bravo, et al., 2016; Sedgmond, et al., 2019). Our data supports the robustness of healthy eating behaviours against perturbation in cortical activity 510 511 within the DLPFC, which is assumed to occur in populations that are obese or with BED 512 (Boeka & Lokken, 2011; Karhunen, et al., 2000). We therefore speculate that there is a 513 diminishing return for attempting to increase neuronal activity within the DLPFC when 514 participants are already able to control their eating behaviours. Hyper-activity in this cortical 515 region may have a ceiling effect beyond which no further improvement is seen. This may 516 account for the null effect we found for food craving, reward and appetite following tDCS,

and can be supported by the moderate-to-strong evidence in favour of the null hypothesis ashighlighted by our Bayesian statistical approach.

519

520 The present study is not without limitation. It is understood that males and females 521 experience different eating behaviours, and may express differing behavioural traits (Rolls, 522 Fedoroff, & Guthrie, 1991). The present study recruited both male and female participants, 523 which may have influenced the effects of tDCS, and provided an additional source of 524 variation across data. However, this is not without precedent as prior studies that recruited 525 male and female participants have shown an experimental effect (Burgess, et al., 2016; 526 Carvalho, et al., 2019; Goldman, et al., 2011). Given the novelty of using the LFPQ it was 527 important to consider the wider effects of tDCS on this variable before focussing on specific 528 sex. Second, our original hypotheses did not consider the impact of eating behaviour traits 529 and as such these were not controlled for during screening. Our inclusion criteria focussed 530 on weight status, but the participants recruited displayed behaviour traits that do not suggest 531 susceptibility to overconsumption, as discussed above; notably, all participants scored below 532 the 50-point cut-off for trait food craving. Third, prior studies have induced hyperactivity in 533 the DLPFC through bilateral and unilateral stimulation of the cortex (Carvalho, et al., 2019; 534 Fassini, et al., 2019; Lapenta, et al., 2014; Ljubisavljevic, et al., 2016). Although these 535 montages have been shown to improve measures of hedonic appetite (Fregni, et al., 2008; 536 Goldman, et al., 2011), the efficacy of such electrode placement has been debated due to 537 the simultaneous effects of anodal and cathodal stimulation on the same cortical region 538 (Bestmann, et al., 2015). The inhibitory effects associated with cathodal stimulation during 539 traditional montages may also impact hedonic appetite measures, as the left DLPFC is 540 implicated in dietary control and food choice behaviour (Higuera-Hernández, et al., 2018). 541 Similar to the right DLPFC, there is some support for reduced activity in the left DLPFC in 542 response to food, when comparing individuals who are lean with those who are obese (Le, et 543 al., 2006; Le, et al., 2007). In the present study, a prefrontal-occipital montage was used, 544 utilising a similar montage seen in previous work (Marron, et al., 2018; Vitor-Costa, et al.,

545 2015). The ability of this montage to induce hyperactivity in the DLPFC has been confirmed 546 in several recent computational models (Marron, et al., 2018; Zheng, et al., 2016; Zheng, et 547 al., 2017). Moreover, we verified that the electric current was being delivered in a consistent 548 manner across all 42 stimulation sessions by checking impedance. Finally, the effectiveness 549 of common sham procedures as a blinding technique has been debated due to significantly 550 greater sensations often reported following active tDCS (Horvath, 2015). Indeed, in the 551 present study participants reported significantly greater itching, tingling and sleepiness 552 following active stimulation. However, the inability of participants to identify the active 553 protocol beyond the level of chance, despite these heightened sensations, provides further 554 support for the use of standardised sham protocols as a blinding technique in tDCS research 555 (Ambrus, et al., 2012).

556

557 **5. CONCLUSION**

558 Our study is the first to report the effects of tDCS on components of food reward in sample of 559 healthy individuals with no susceptibility to overconsume, and we show no significant change 560 in these measures. Prior to examining these effects, we established an indication of data 561 reliability and revealed some plausible targets for future effects through tDCS exposure. In 562 the present sample these effects were transient for the most part and, in line with the work 563 by Burgess and colleagues (Burgess, et al., 2016; Ray, et al., 2017), this highlights a 564 behaviour trait-dependent effect of stimulation. Future work should focus on populations who 565 are at risk of reward-driven overconsumption and weight gain, such as those showing 566 recurrent binge eating behaviours, as these individuals may be responsive to the effects of 567 tDCS on hedonic appetite.

568

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574 AUTHOR CONTRIBUTIONS

- 575 Jordan Beaumont: Conceptualisation, Methodology, Validation, Formal analysis,
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- 584 Martin Barwood: Conceptualisation, Methodology, Validation, Resources, Writing original
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- 586

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